

Fundamentals of Photorealistic Rendering (Optimizations)

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GIST

Efficiency of MC estimators

- An efficiency measure for estimators F
 - $e(F) = \frac{1}{V(F)T(F)}$
 - To improve efficiency, we need to reduce the variance and time
 - $V(F)$ may be replaced with $\text{MSE}(F)$ for biased estimators

Outline

- Russian Roulette and Splitting
- Importance Sampling
- Multiple Importance Sampling
- Adaptive Rendering (Sampling and Reconstruction)
- Correlated Sampling (Antithetic Variates, Common Random Numbers, Control Variates)
- Gradient-Domain Rendering

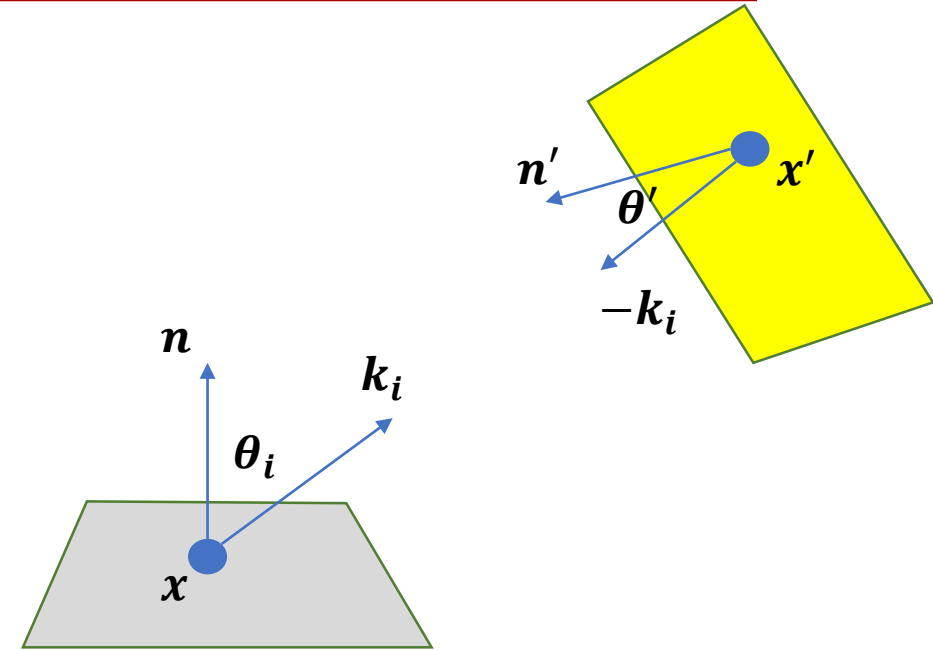
Russian Roulette

- e.g. direct lighting

$$\begin{aligned} \circ L_S(x, k_o) &= \int_{\text{all } x'} \frac{\rho(k_i, k_o) L_e(x', -k_i) v(x, x') \cos\theta_i \cos\theta'}{\|x - x'\|^2} dA' \\ &\approx \frac{1}{N} \sum_{i=1}^N \frac{\rho(k_i, k_o) L_e(x', -k_i) v(x, x') \cos\theta_i \cos\theta' A}{p(x') \|x - x'\|^2} \end{aligned}$$

- Problems

- Expensive to compute $v(x, x')$
- Let suppose that there are some directions k_i where the integrand's value is almost 0 due to $\rho(k_i, k_o) \approx 0$
 - In this case, evaluating $v(x, x')$ is not a good idea, as it decreases the efficiency
- Q. Can we somehow skip these directions while maintaining a correct answer?



Russian Roulette

- Given an estimator F , a new estimator F' with Russian Roulette can be given:

- $F' = \begin{cases} \frac{F-qc}{1-q} & \xi > q \\ c & \text{otherwise} \end{cases}$

- q : termination probability

- c is usually chosen as 0

- Consistency check

- $E(F') = (1 - q) \left(\frac{E(F) - qc}{1 - q} \right) + qc = E(F)$

- Properties

- It does not reduce variance, but improves efficiency by skipping unimportant parts

Russian Roulette

- Examples

- $L_S(k_o) \approx L_e(k_o) + \frac{\rho(k_i, k_o) L_f(k_i) \cos\theta_i d\sigma_i}{p(k_i)}$

- Problem: this is a recursive form, so ray depth can be infinite

- Can apply the Russian Roulette to path tracing so that the ray depth can be reduced

- $F' = \begin{cases} \frac{F-qc}{1-q} & \xi > q \\ c & \text{otherwise} \end{cases}$

- Given a termination probability $q = 0.5$ and $c = 0$, we can terminate the reflection with the probability. In this case, we need to scale the radiance value with $\frac{1}{1-q}$

Russian Roulette

- Bad examples
 - Apply this to the camera rays with a termination probability $q = 0.99$
 - In this case,
 - Only trace 1% of the camera rays
 - Most of pixels are black and a few pixels are very bright, although its expectation is still correct
 - The variance of the estimator will be much higher than the original estimator

Splitting

- Splitting is a technique to increase the number of samples for improving the efficiency
 - Allocate more rays to important dimensions

- E.g., direct lighting with a shortened version

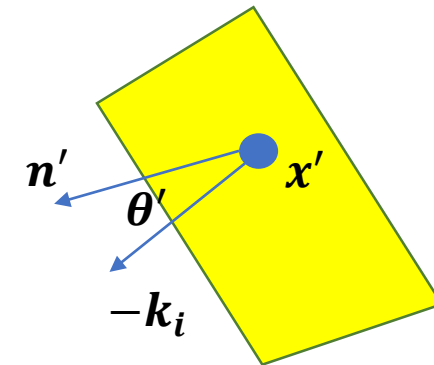
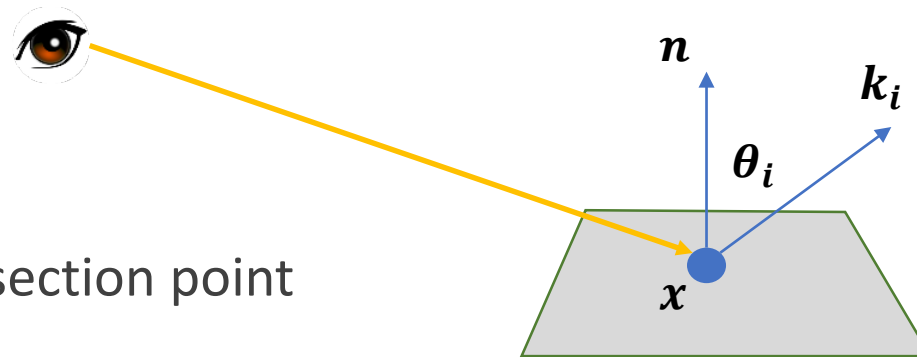
- $\int_A \int_S L_d(x, y, w) dx dy dw$

- A: pixel area

- S: light area

- L_d : exitant radiance at the intersection point

- (x, y) : position on image



Splitting

- direct lighting with a shortened version

- $\int_A \int_S L_d(x, y, w) dx dy dw$

- A straightforward approach:

- Generate N samples $(x_1, y_1, w_1), \dots, (x_N, y_N, w_N)$

- Evaluate $L_d(x_1, y_1, w_1), \dots, L_d(x_N, y_N, w_N)$

- Need to generate N shadow rays

- Average the radiance values

- Typically need a lot of samples given a large area light or many point lights

- e.g., N = 100, 200 rays will be used (100 primary rays, 100 shadow rays)

- Issue: 100 primary rays are often too much for a good antialiasing. Can we focus on

Splitting

- Typically need a lot of samples given a large area light or many point lights
 - e.g., $N = 100$, 200 rays will be used (100 primary rays, 100 shadow rays)
 - Issue: 100 primary rays are often too many for a good antialiasing.
 - Splitting
 - $\frac{1}{N} \frac{1}{M} \sum_{i=1}^N \sum_{j=1}^M \frac{L(x_i, y_i, w_{i,j})}{p(x_i, y_i) p(w_{i,j})}$
 - Are able to use 5 image samples and take 20 light samples per image sample. Total ray # = $5 + 5 \times 20 = 105$
 - Still use 100 shadow rays for high-quality soft shadows, but the total number of rays is reduced

How to optimize the parameters?

- Tuning the parameters in Russian roulette and splitting is often done empirically
- Automatic control?
 - EARS: efficiency-aware russian roulette and splitting, Rath et al. 2022

Outline

- Russian Roulette and Splitting
- **Importance Sampling**
- Multiple Importance Sampling
- Adaptive Rendering (Sampling and Reconstruction)
- Correlated Sampling (Antithetic Variates, Common Random Numbers, Control Variates)
- Gradient-Domain Rendering

Importance Sampling

- A variance reduction technique for Monte Carlo estimators
 - Allocate more samples to the important region where the integrand's value is high
- A Monte Carlo estimator:
 - $F_N = \frac{1}{N} \sum_{i=1}^N \frac{f(X_i)}{p(X_i)}$
- e.g., $L_S(k_o) \approx L_e(k_o) + \frac{\rho(k_i, k_o) L_f(k_i) \cos \theta_i d\sigma_i}{p(k_i)}$
 - An intuition is that if we take more samples in terms of k_i that makes $\cos \theta_i$ high, we can reduce the variance of the estimator

Importance Sampling

- e.g., Evaluate an integral $\int f(x)dx$
- Note that we can choose an arbitrary pdf, $p(x)$
- What if we choose $p(x) \propto f(x)$ or $p(x) = cf(x)$
 - $c = \frac{1}{\int f(x)dx}$ a constant for normalization
 - $F_N = \frac{1}{N} \sum_{i=1}^N \frac{f(X_i)}{p(X_i)} = \frac{1}{N} \sum_{i=1}^N \frac{1}{c}$
 - $V(F_N) = 0$
- Notes
 - In practice, we cannot choose $p(x)$ in this way, but it provides some intuition
 - If we choose $p(x)$ similarly compared to the shape of $f(x)$, we are able to reduce variance

Importance Sampling

- A counterexample:
 - Evaluate an integral $\int f(x)dx$
 - $p(x) = \begin{cases} 99.01 & x \in [0,0.01) \\ 0.01 & x \in [0.01,1] \end{cases}$
 - $f(x) = \begin{cases} 0.01 & x \in [0,0.01) \\ 1.01 & x \in [0.01,1] \end{cases}$
 - $\int f(x)dx = 1$
 - Most of samples will be taken from $[0,0.01)$, and $\frac{f(x_i)}{p(x_i)} \approx 0.0001$ which is far from 1
 - In this case, the variance will increase a lot
- Note
 - In practice, it is easy to apply an important sampling to the rendering, by considering only some terms
 - Taking account for all terms is ideal but this can be very challenging

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Multiple Importance Sampling

- Evaluate an integral $\int f(x)g(x)dx$
- If we have each important sampling scheme for the functions $f(x)$ and $g(x)$, how can we combine the techniques?
- Multiple important sampling addresses this issue and this will be discussed later

Multiple Importance Sampling

- Proposed by Eric Veach
- Assume we have multiple (more than one) sampling techniques
- Q. How do we combine the techniques?
- Motivation:
 - Light transport integral is complex (most terms are unknown and should be estimated)
 - Designing a sampling technique, which works well for a variety of situations, is difficult

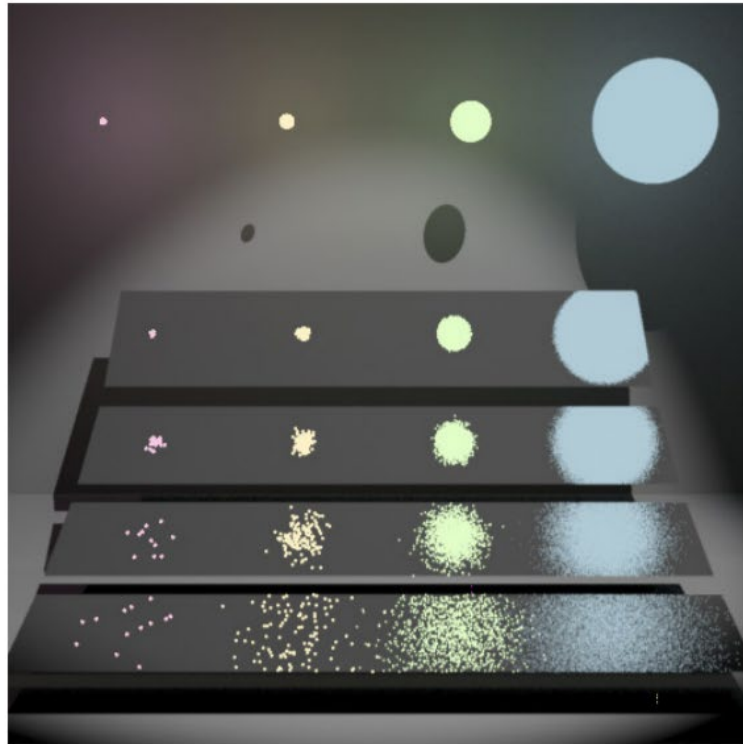
Applications

- Glossy highlights from area light sources
- Common sampling techniques
 - Sampling the light sources
 - Sampling the BRDF

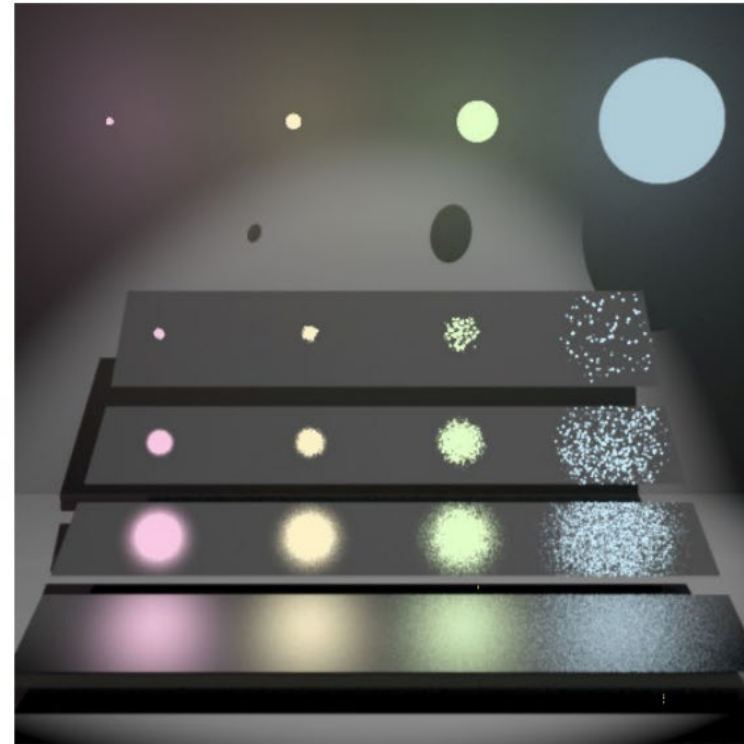
- Light transport for direct lighting

$$\circ L_S(x, k_o) = \int_{all\ x'} \frac{\rho(k_i, k_o) L_e(x', -k_i) v(x, x') \cos\theta_i \cos\theta'}{\|x - x'\|^2} dA'$$

Applications



(a) Sampling the BSDF



(b) Sampling the light sources

Applications

- Sampling the light sources

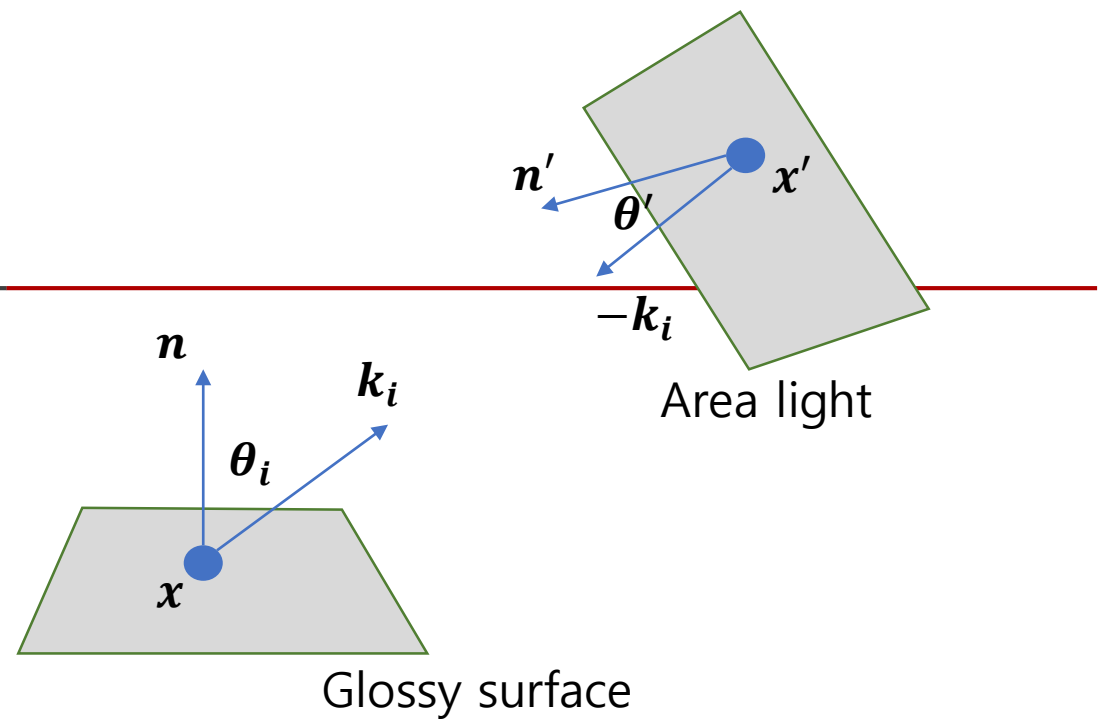
$$\circ p(x') \propto \frac{L_e(x', -k_i) \cos\theta_i \cos\theta'}{\|x-x'\|^2}$$

- Sampling the BRDF

$$\circ p(k_i) \propto \rho(k_i, k_o)$$

- Light transport for direct lighting

$$\circ L_S(x, k_o) = \int_{\text{all } x'} \frac{\rho(k_i, k_o) L_e(x', -k_i) v(x, x') \cos\theta_i \cos\theta'}{\|x-x'\|^2} dA'$$



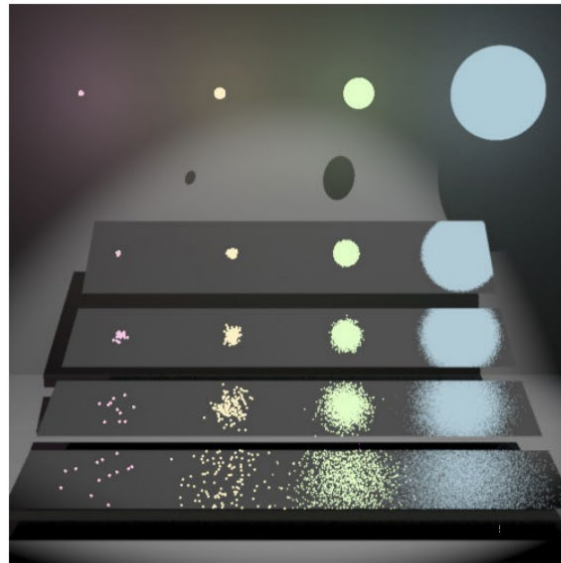
Multi-Sample Estimator

- A combination strategy to average samples from multiple sampling techniques
- $\int_{\Omega} f(x) d\mu(x)$
 - $f : \Omega \rightarrow R$
- Samples ($j = 1, \dots, n_i$) from i-th sampling
 - $X_{i,j}$
- Multi-sample estimator
 - $F = \sum_{i=1}^n 1/n_i \sum_{j=1}^{n_i} w_i(X_{i,j}) \frac{f(X_{i,j})}{p_i(X_{i,j})}$
 - Conditions for unbiasedness
 - $\sum_{i=1}^n w_i(x) = 1$ whenever $f(x) \neq 0$
 - $w_i(x) = 0$ whenever $p_i(x) = 0$
 - See the Veach's thesis for the proof.

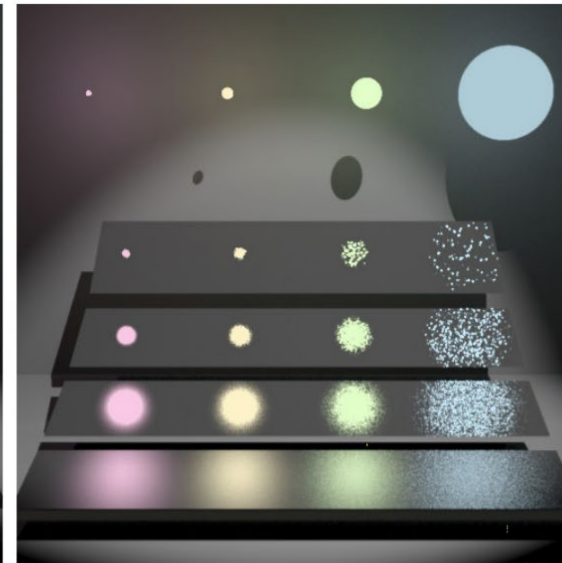
Multi-Sample Estimator

- The balance heuristic

$$\circ \hat{w}_i(x) = \frac{n_i p_i(x)}{\sum_k n_k p_k(x)}$$



(a) Sampling the BSDF



(b) Sampling the light sources



Images from the Veach's thesis

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Background

- Monte Carlo estimator

- $\hat{\theta} = \frac{1}{N} \sum_i f(X_i)$

- $var(\hat{\theta}) = \frac{\sigma^2}{N}$

- Homogeneous error

- $var(f(X_i)) = \sigma^2$

- Uniform sampling rate would make sense for the homogeneous error

- Heterogeneous error

- Variance can vary across local regions

- Light transport function is typically nonlinear and variance varies across regions

Classical Adaptive Sampling

- Image-space adaptive sampling using per-pixel variance
 - Per-pixel estimator
 - Imagine you have an unknown image-space function per pixel
 - $\widehat{\theta}_p = \frac{1}{N_p} \sum_i f_p(X_i)$
 - $\text{var}(\widehat{\theta}_p) = \frac{\sigma_p^2}{N_p}$
 - Assumption:
 - Random samples $f_p(X_i)$ are iid and share a variance σ_p^2
 - Note that σ_p^2 can vary per pixel p
 - A simple approach
 - Control the number of samples per pixel N_p to be proportional to σ_p^2
 - High-error regions have high sample counts
 - How do we estimate σ_p^2 ?
 - $s_p^2 = \frac{1}{N_p-1} \sum_i (f_p(X_i) - \bar{f}_p(X))^2$

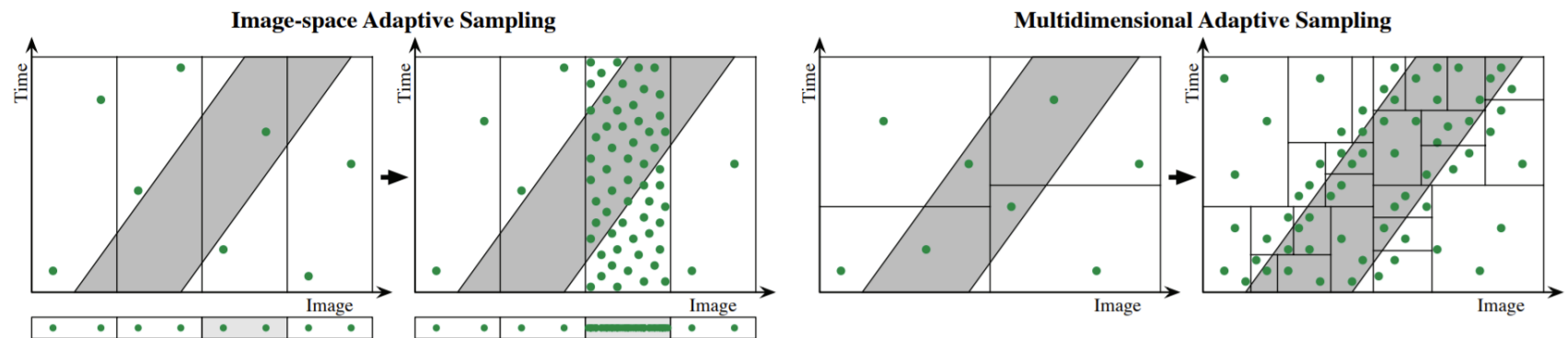
Classical Adaptive Sampling

- Two-step approach using the sample variance
 - Allocate a small number of samples (equal sample count) per pixel
 - Estimate the sample variance and decide the number of samples per pixel for the next stage
 - Repeat until we use a target sample count

- Potential problem of the simple approach
 - Per-pixel sample variances can be typically noisy unless a large number of samples are used.
 - Sample counts can be noisy and it typically leads to a noisy rendered image
 - A simple trick
 - Apply an image filter to the sample variance and use the smoothed variance

Multi-dimensional Adaptive Sampling

- The classical approach only controls the number of samples in image space
 - Sampling density for the other dimensions (lens, time, ...) is still uniform.
- Multi-dimensional sampling
 - The random sample is in a high-dimensional space
 - Maintain a high-dimensional structure (kd-trees) that stores the samples
 - This idea was proposed initially in [Kajiya 86], but it was fully accomplished in [Hachisuka et al. 08]



Multi-dimensional Adaptive Sampling

- Iteratively allocate more samples to the kd-tree leaves whose variances are high
- Kd-tree leafs are subdivided if its number of samples are higher than a user-defined threshold
- The final image is reconstructed by projecting the sub-regions to the image plane (sub-regions are assumed to be constant)

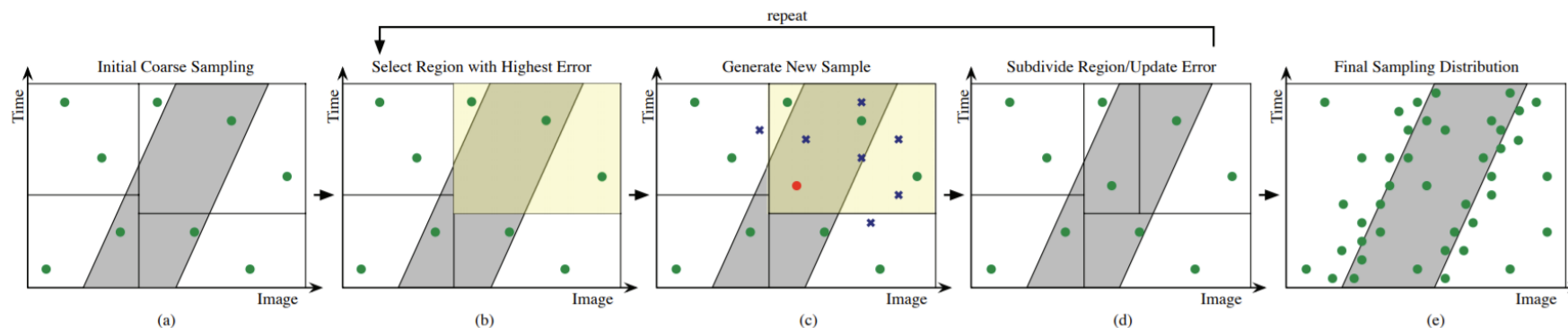


Image from [Hachisuka et al. 08]

Multi-dimensional Adaptive Sampling

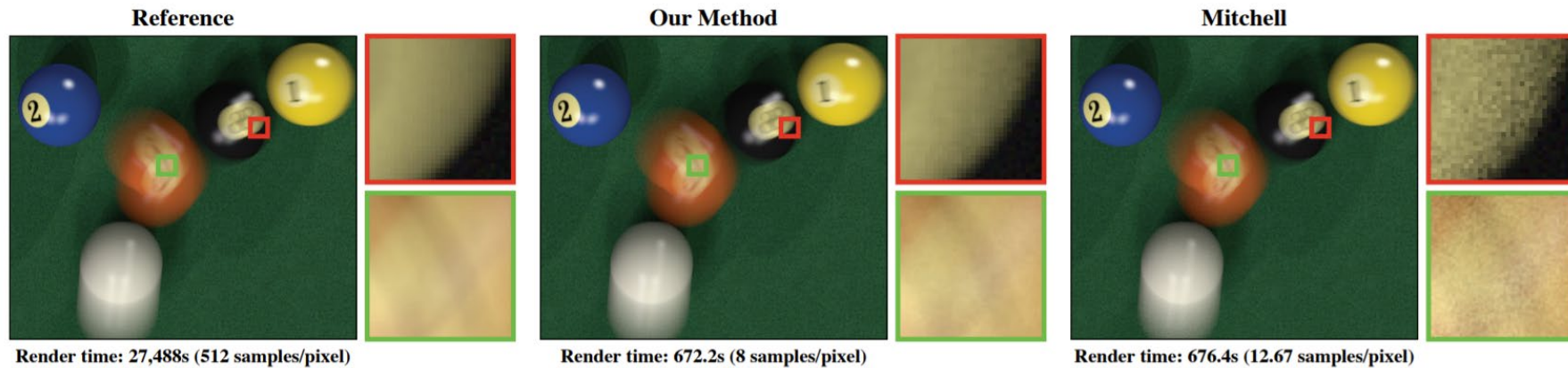
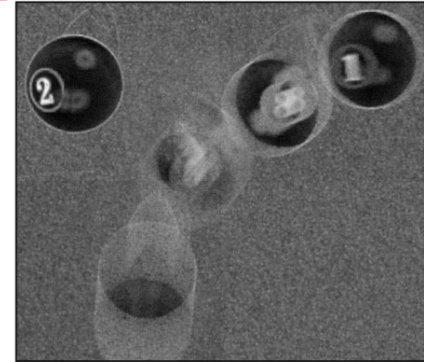


Image from [Hachisuka et al. 08]

Multi-dimensional Adaptive Sampling

- Pros.
 - Very effective for a moderate dimension (direct illumination, motion blur, and depth-of-fields)
- Cons.
 - Sample density becomes very sparse as increasing the dimensionality of samples (e.g., for global illumination)

Adaptive Sampling using Wavelet Space

- Adaptive Wavelet Rendering [Overbeck et al. 09]
- Properties
 - Image-space adaptive sampling
 - More robust than the classical approach
- Wavelet thresholding using variances

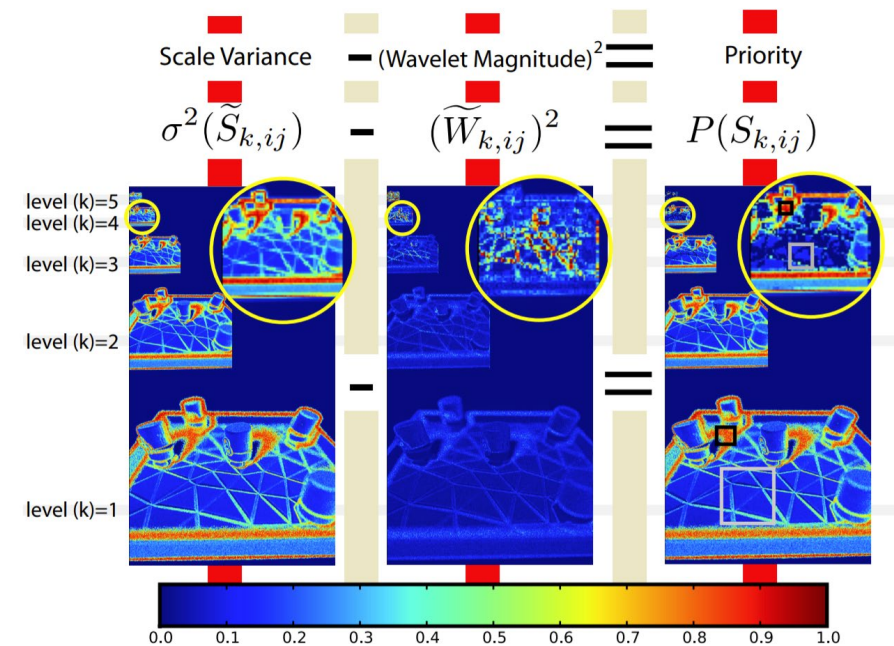
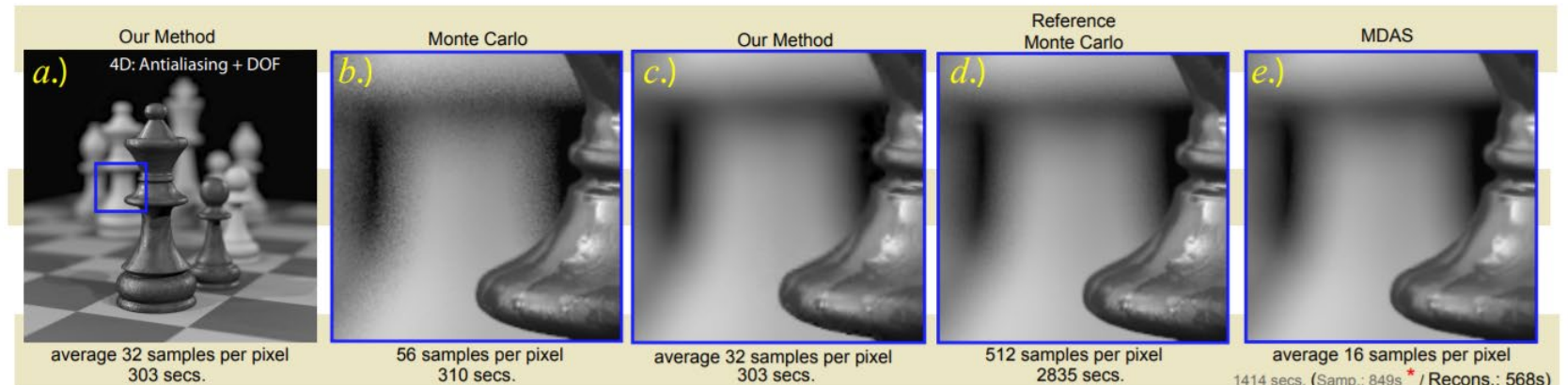


Image from [Overbeck et al. 09]

Adaptive Sampling using Wavelet Space

- Pros.
 - Works well for depth-of-field or motion blur
- Cons.
 - Produces ringing artifacts when GI simulates



Classical Image Space Adaptive Rendering

- Controlling sample density using sample variances is not optimal
- Suppose a typical combination:
 - Path tracing + image denoising
 - The variance of the final image (reconstruction variance) can be very different from the sample variance
 - Adaptive sampling should allocate more samples to the regions with high reconstruction errors

Classical Image Space Adaptive Rendering

- Adaptive Sampling and Reconstruction using Greedy Error Minimization
 - [Rousselle et al. 11]
- Control two parameters
 - Sampling rate per pixel
 - Reconstruction parameter per pixel
- Target combination
 - Path tracing + Gaussian filter

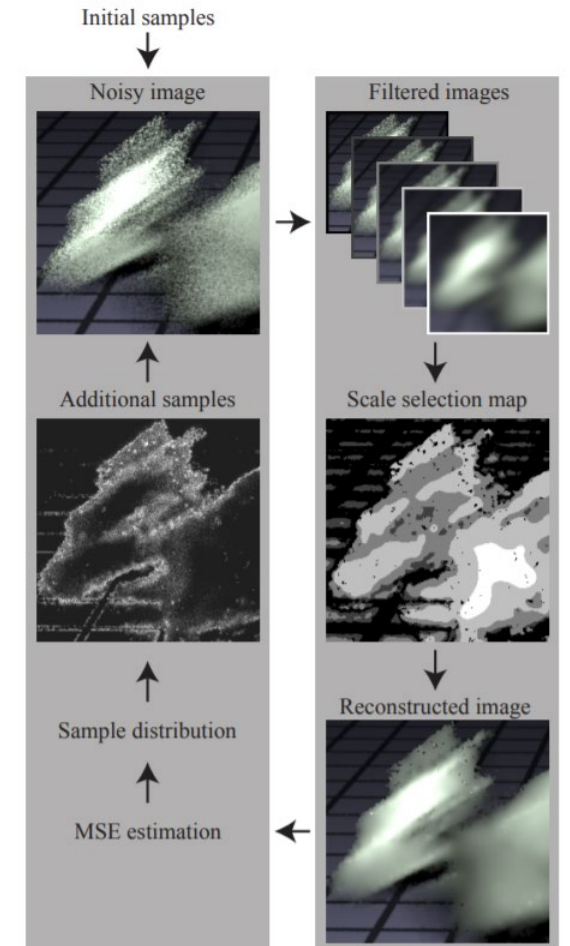


Image from [Rousselle et al. 11]

Classical Image Space Adaptive Rendering

- Adaptive Sampling and Reconstruction using Greedy Error Minimization
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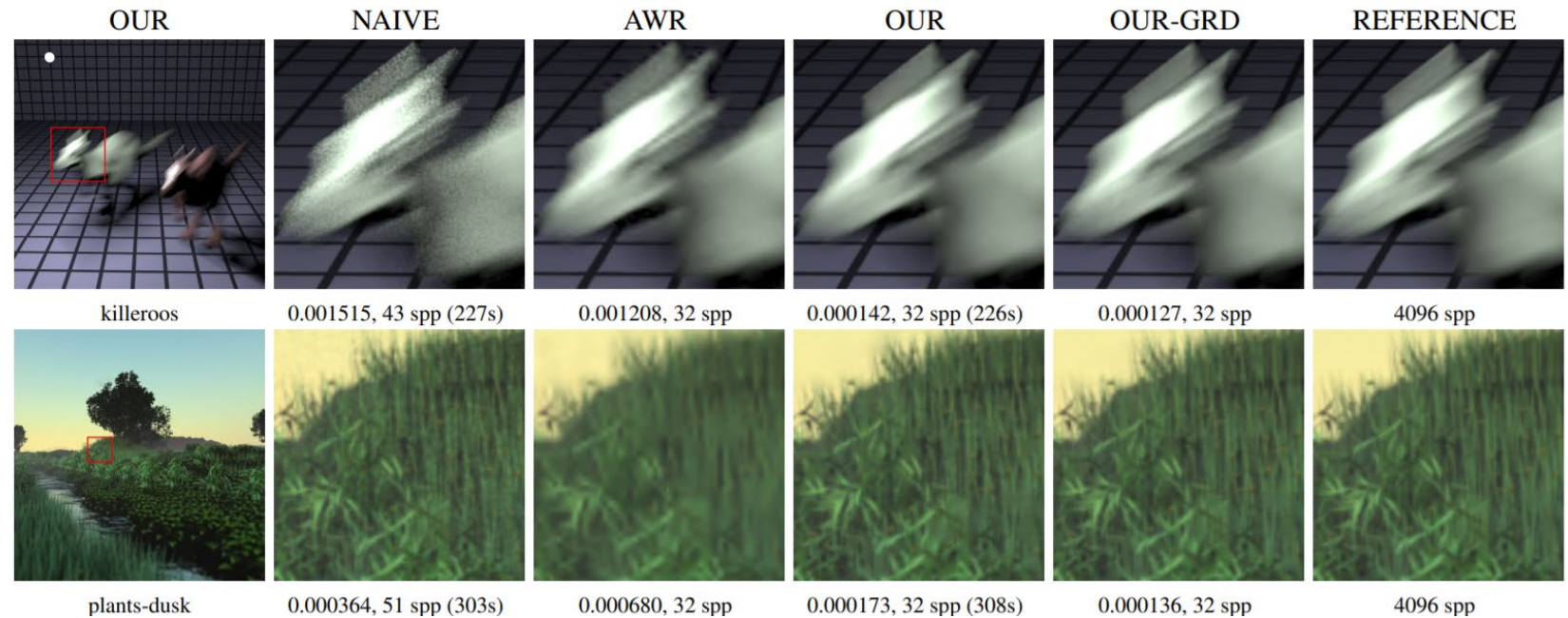


Image from [Rousselle et al. 11]

Classical Image Space Adaptive Rendering

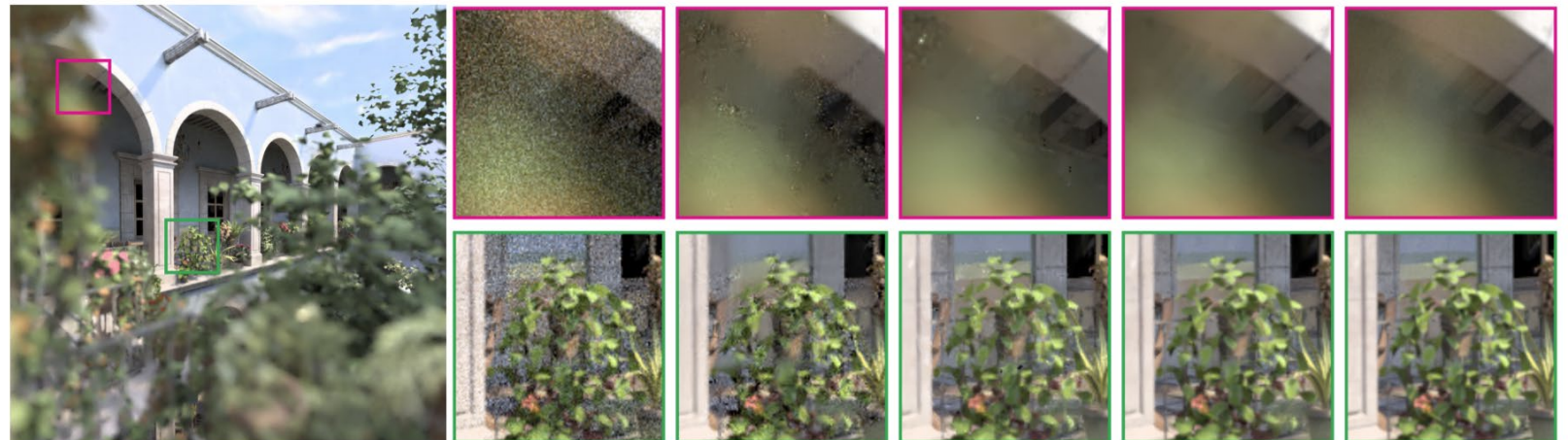
- Adaptive Sampling and Reconstruction using Greedy Error Minimization
 - [Rousselle et al. 11]
- Pros.
 - Very simple and fast
 - Work reasonably well
- Cons.
 - Isotropic filter (Gaussian) is not effective for non-linear functions (image edges)

Classical Image Space Adaptive Rendering

- Nonlinear reconstructions
 - Adaptive Rendering with Non-Local Means Filtering [Rousselle et al. 12]
 - Non-local means
 - On Filtering the Noise from the Random Parameters in Monte Carlo Rendering [Sen and Darabi 11]
 - Cross bilateral filtering
 - Adaptive Rendering based on Weighted Local Regression [Moon et al. 14]
 - Local regression
 - Nonlinearly Weighted First-order Regression for Denoising Monte Carlo Renderings [Bitterli et al. 16]
 - Local regression + non-local means

Classical Image Space Adaptive Rendering

- General procedure
 - Allocate sparse samples initially
 - Reconstruct the image using a non-linear filtering
 - Estimate optimal parameters for the reconstruction
 - Estimate MSE (bias² + variance) of the image
 - Allocate more samples to the image regions with high MSEs



Classical Image Space Adaptive Rendering

- Reconstruction performance depends on:
 - Nonlinearity of the chosen filter
 - Robust estimation of errors
 - Rendering-specific features to identify high-frequency image information (edges)
- Adaptive sampling is closely related to a chosen reconstruction method

Neural Image Space Adaptive Rendering

- So far:
 - Reconstruction filters are hand-crafted ones
 - Estimating MSE is often a tedious task (and also quite difficult to be robust)

- Deep learning based approaches
 - Kernel-Predicting Convolutional Networks for Denoising Monte Carlo Renderings [Bako et al. 17]
 - A general weighting can be represented as
 - $\hat{y}_c = \sum_i w_i y_i$
 - $w_i > 0$

 - Key advantage is that we don't need to assume a specific functional form for w_i

Neural Image Space Adaptive Rendering

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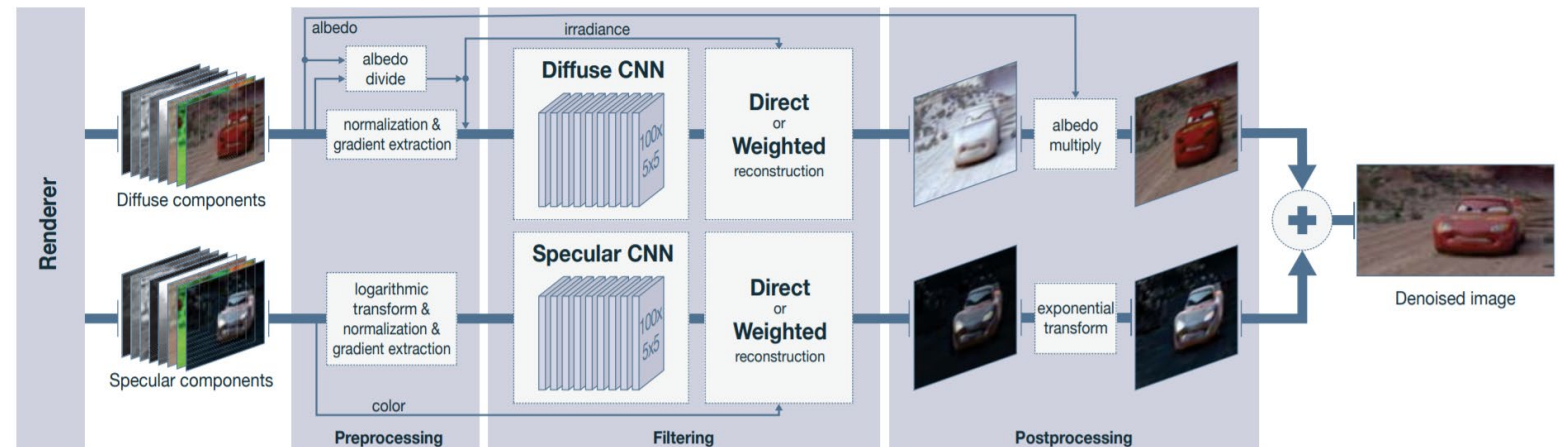


Image from [Bako et al. 17]

Neural Image Space Adaptive Rendering

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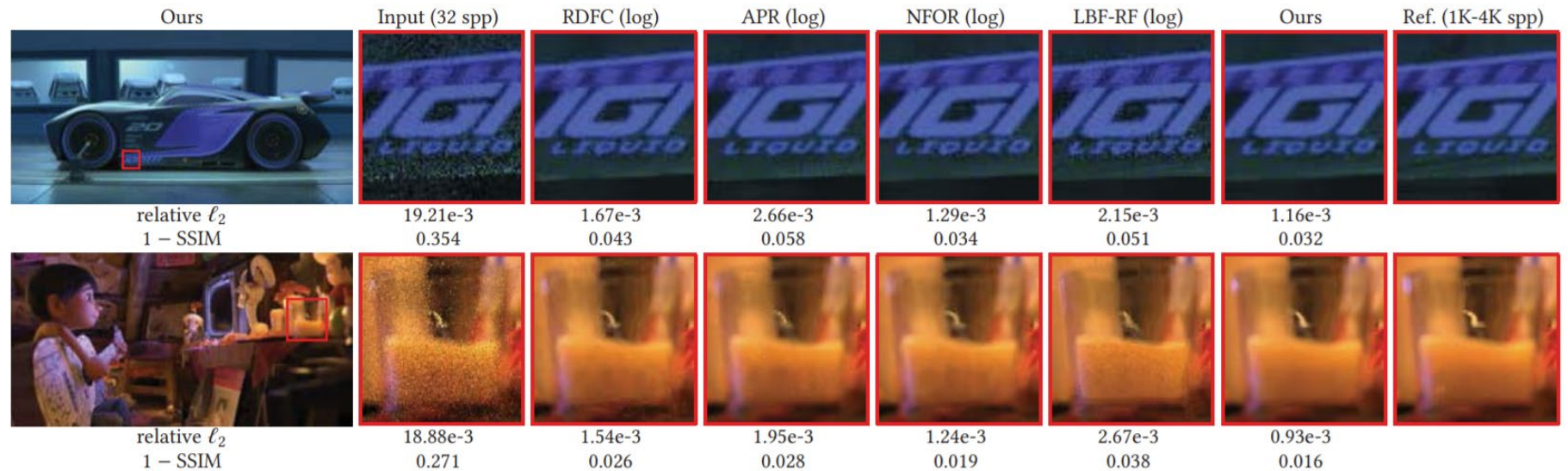
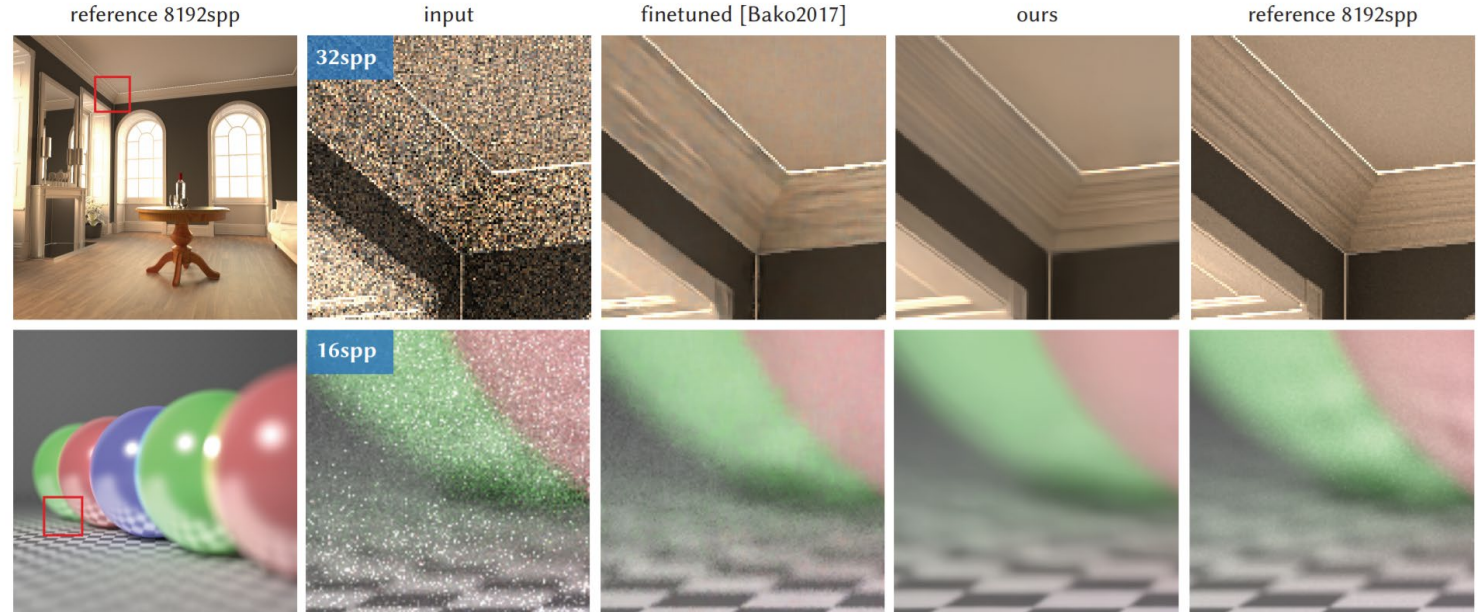


Image from [Bako et al. 17]

Neural Image Space Adaptive Rendering

- Deep learning based approaches
 - Sample-based Monte Carlo Denoising using a Kernel-Splatting Network [Gharbi et al. 19]
 - A general weighting in a sample space can be represented as

- $\hat{y}_{c,j} = \sum_i \sum_j w_{i,j} y_{i,j}$



Recent Optimizations for Adaptive Rendering

- Larger receptive fields with advanced NNs
 - Monte Carlo Denoising via Auxiliary Feature Guided Self-Attention, Yu et al. 2021
- Post-correction via a neural network trained only using a test image (without any pretraining)
 - Self-Supervised Post-Correction for Monte Carlo Denoising, Back et al. 2022

Recent Optimizations for Adaptive Rendering

- Reconstruction bias is not directly controllable by neural image reconstruction
 - i.e., not consistent: Simply adding more samples does not guarantee error reduction
- Making an image denoiser consistent
 - Neural James-Stein Combiner for Unbiased and Biased Renderings, Gu et al. 2022

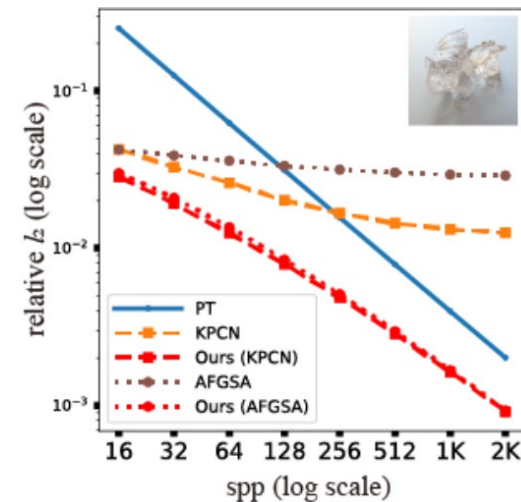


Image from [Gu et al. 2022]

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Antithetic Variates

- $\theta = \int_{\Omega} f(x)d\mu(x)$
- Monte Carlo estimator
 - $\hat{\theta} = \frac{1}{N} \sum_i f(X_i)$
- Suppose:
 - $pdf(X) = pdf(-X)$
 - i.e., a symmetric pdf of X
- Define a variable
 - $Y_i \equiv \frac{f(X_i) + f(-X_i)}{2}$
 - Y_i is an unbiased estimate of θ , $E[Y] = E[f(X)] = \theta$
 - $\hat{\theta}^{av} = \frac{1}{N} \sum_i Y_i$

Antithetic Variates

- (original) Monte Carlo estimator

- $\hat{\theta} = \frac{1}{N} \sum_i f(X_i)$

- New estimator

- $\hat{\theta}^{av} = \frac{1}{N} \sum_i Y_i$

- Suppose i.i.D:

- $\text{var}(f(X_i)) = \text{var}(f(X_j)) = \sigma^2$

- Variance of estimators

- $\text{var}(\hat{\theta}) = \text{var}\left(\frac{1}{N} \sum_i f(X_i)\right) = \frac{1}{N^2} \sum_i \text{var}(f(X_i)) = \frac{1}{N^2} \sum_i \sigma^2 = \frac{\sigma^2}{N}$

- $\text{var}(\hat{\theta}^{av}) = \text{var}\left(\frac{1}{N} \sum_i Y_i\right) = \frac{1}{N^2} \sum_i \text{var}(Y_i)$

Antithetic Variates

- Variance of estimators

- $var(\hat{\theta}) = var\left(\frac{1}{N}\sum_i f(X_i)\right) = \frac{1}{N^2}\sum_i var(f(X_i)) = \frac{1}{N^2}\sum_i \sigma^2 = \frac{\sigma^2}{N}$

- $var(\hat{\theta}^{av}) = var\left(\frac{1}{N}\sum_i Y_i\right) = \frac{1}{N^2}\sum_i var(Y_i)$

- $var(Y_i) = var\left(\frac{f(X_i)+f(-X_i)}{2}\right) = \frac{1}{4}[var(f(X_i)) + var(f(-X_i)) + 2cov(f(X_i), f(-X_i))]$

- Putting $var(Y_i)$ into $var(\hat{\theta}^{av})$:

- $var(\hat{\theta}^{av}) = \frac{1}{4N^2}\sum_i\{2\sigma^2 + 2cov(f(X_i), f(-X_i))\} = \frac{1}{2N^2}\sum_i\{\sigma^2 + cov(f(X_i), f(-X_i))\}$

- If there is no correlation

- $var(\hat{\theta}^{av}) = \frac{\sigma^2}{2N} = \frac{var(\hat{\theta})}{2}$

- No actual gain here since we use 2N samples instead of N samples

- What if there is a negative correlation?

- $var(\hat{\theta}^{av}) < \frac{var(\hat{\theta})}{2}$

Antithetic Variates

- Antithetic variates introduces a negative correlation for monotonically increasing functions
 - $cov(f(X_i), f(-X_i)) < 0$
 - e.g., linear functions – ideal case
- Properties
 - Very simple to implement it even for high-dimensional cases
 - Some applications in rendering:
 - Direct lighting
 - Pixel estimator in PSS?

Common Random Numbers (CRN)

- Suppose we want to estimate a difference between two functions

- $\theta_1 = \int_{\Omega} f_1(x) d\mu(x)$

- $\theta_2 = \int_{\Omega} f_2(x) d\mu(x)$

- $\theta = \theta_1 - \theta_2$

- MC estimator

- $\hat{\theta} = \frac{1}{N} \sum_i f_1(X_i) - \frac{1}{N} \sum_j f_2(X_j)$

- CRN estimator

- $\hat{\theta} = \frac{1}{N} \sum_i f_1(X_i) - \frac{1}{N} \sum_i f_2(X_i) = \frac{1}{N} \sum_i (f_1(X_i) - f_2(X_i))$

Common Random Numbers (CRN)

- CRN estimator

- $\hat{\theta} = \frac{1}{N} \sum_i f_1(X_i) - \frac{1}{N} \sum_i f_2(X_i) = \frac{1}{N} \sum_i (f_1(X_i) - f_2(X_i))$

- $var(\hat{\theta}) = var\left(\frac{1}{N} \sum_i (f_1(X_i) - f_2(X_i))\right) = 1/N^2 \sum_i var(f_1(X_i) - f_2(X_i))$

- $var(f_1(X_i) - f_2(X_i)) = var(f_1(X_i)) + var(f_2(X_i)) - 2cov(f_1(X_i), f_2(X_i))$

- What if $cov(f_1(X_i), f_2(X_i)) = 0$?

- No actual gain over the ordinary MC estimator.

Common Random Numbers (CRN)

- When the two functions tend to increase (or decrease) together,
 - $cov(f_1(X_i), f_2(X_i)) > 0$
 - e.g., both functions are linear whose derivatives have the same sign.
- Applications in rendering
 - Estimating image gradients
 - (screened) Poisson reconstruction takes the image gradients to output a reconstructed image
- Q. can we decide whether or not we apply the CRN?
 - In practice, it is hard to know if there is such correlation in advance.
 - However, implementing and testing CRN are very easy.

CRN examples



Path tracing with CRN numbers, 76 samples per pixel

Control Variates

- $\theta = \int_{\Omega} f(x)d\mu(x)$
- Monte Carlo estimator
 - $\hat{\theta} = \frac{1}{N} \sum_i f(X_i)$
- Define a control variate $g(x)$ whose integration G is known.
 - $\theta = \int_{\Omega} f(x) - \alpha g(x)d\mu(x) + \alpha G$
 - $\hat{\theta}^{cv} = \frac{1}{N} \sum_i (f(X_i) - \alpha g(X_i)) + \alpha G$

Control Variates

- When $\alpha = 1$
 - $\hat{\theta}^{cv} = \frac{1}{N} \sum_i (f(X_i) - g(X_i)) + G$
 - $var(\hat{\theta}^{cv}) = \frac{1}{N^2} \sum_i var(f(X_i) - g(X_i)) = 1/N^2 \{ \sum_i var(f(X_i)) + var(g(X_i)) - 2cov(f(X_i), g(X_i)) \}$
 - Assume that:
 - $var(f(X_i)) = var(g(X_i)) = \sigma^2$
 - $cov(f(X_i), g(X_i)) = \sigma_{f,g}^2$
 - $corr(f(X_i), g(X_i)) = corr_{f,g} = \frac{\sigma_{f,g}^2}{\sigma_f \sigma_g} = \sigma_{f,g}^2 / \sigma^2$
 - $var(\hat{\theta}^{cv}) = \frac{\sum_i 2\sigma^2}{N^2} - \frac{2\sum_i \sigma_{f,g}^2}{N^2} = \frac{2\sigma^2}{N} - \frac{2\sigma_{f,g}^2}{N}$
 - Variance of the original estimator with the same sample count N
 - $var(\hat{\theta}) = \frac{\sigma^2}{2N}$
 - Condition for $var(\hat{\theta}^{cv}) < var(\hat{\theta})$
 - $\frac{2\sigma^2}{N} - \frac{2\sigma_{f,g}^2}{N} < \frac{\sigma^2}{2N}$
 - $\frac{3\sigma^2}{4} < \sigma_{f,g}^2$
 - $\frac{3}{4} < corr_{f,g}$

Control Variates

- When $\alpha = \frac{\sigma_{f,g}^2}{\sigma_g^2} = \frac{\sigma_{f,g}^2}{\sigma^2}$
 - $\hat{\theta}^{cv} = \frac{1}{N} \sum_i (f(X_i) - \alpha g(X_i)) + \alpha G$
 - $var(\hat{\theta}^{cv}) = \frac{1}{N^2} \sum_i var(f(X_i) - \alpha g(X_i)) = 1/N^2 \{ \sum_i var(f(X_i)) + \alpha^2 var(g(X_i)) - 2\alpha cov(f(X_i), g(X_i)) \}$
 - $= \frac{\sigma^2}{N} + \frac{\alpha^2 \sigma^2}{N} - \frac{2\alpha \sigma_{f,g}^2}{N}$
 - $= \frac{\sigma^2}{N} + \frac{corr_{f,g}^2 \sigma^2}{N} - \frac{2corr_{f,g}^2 \sigma^2}{N}$
 - $= \frac{\sigma^2}{N} (1 - corr_{f,g}^2)$
 - Condition for $var(\hat{\theta}^{cv}) < var(\hat{\theta})$
 - $\frac{\sigma^2}{N} (1 - corr_{f,g}^2) < \frac{\sigma^2}{2N}$
 - $1/2 < corr_{f,g}^2$
 - $\frac{1}{4} < |corr_{f,g}|$

Outline

- Russian Roulette and Splitting
- Importance Sampling
- Multiple Importance Sampling
- Adaptive Rendering (Sampling and Reconstruction)
- Correlated Sampling (Antithetic Variates, Common Random Numbers, Control Variates)
- Gradient-Domain Rendering

Gradient-Domain Rendering

- Image gradients can be estimated via correlated sampling
 - $I(x + 1, y) - I(x, y)$
 - $I(x, y + 1) - I(x, y)$
- The variance of the estimated gradients can be smaller than the pixel color when then the covariance term is positive
 - $var(I(x + 1, y) - I(x, y)) = var(I(x + 1, y)) + var(I(x, y)) - 2cov(I(x + 1, y), I(x, y))$

Gradient-Domain Rendering

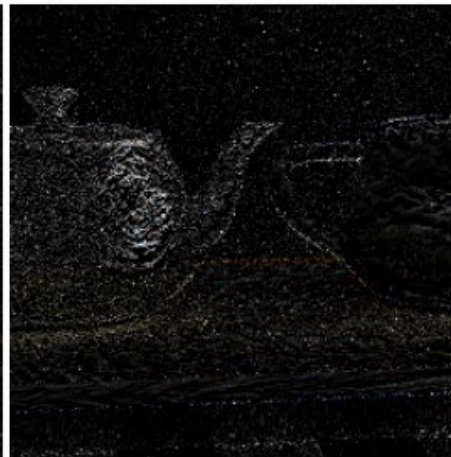
- Rendering estimates three images:
 - Primal colors (e.g., standard path tracing)
 - Image gradients (e.g., correlated samplings such as CRNs, shift mapping, and path reusing)



(a) Input color image, 512 spp
relMSE 0.1318



(b) Input gradients in
horizontal direction



(c) Input gradients in
vertical direction



(d) L2
relMSE 0.0184

Gradient-Domain Rendering

- Screened Poisson Reconstruction

$$\circ \hat{y} = \operatorname{argmin}_{\bar{y}} \sum_{i=1}^n \|\alpha(y_i - \bar{y}_i)\|^2 + \sum_{i=1}^n \|g_i^{\text{dx}} - D^{\text{dx}}\bar{y}_i\|^2 + \sum_{i=1}^n \|g_i^{\text{dy}} - D^{\text{dy}}\bar{y}_i\|^2$$

- $g_i^{\text{dx}}, g_i^{\text{dy}}$: Estimated gradients at pixel i in x and y directions
- y_i : Pixel color at i
- α : user-parameter (e.g., 0.2)
- $D^{\text{dx}}, D^{\text{dy}}$: differential operator in x and y directions (i.e., finite differences)
- Has a closed-form solution (i.e., normal equation) when the norm is L2

Gradient-Domain Rendering

- When L2 reconstruction is used, the output is unbiased.
- One may use a neural network that takes the three inputs
 - e.g., Deep Convolutional Reconstruction for Gradient-Domain Rendering, Kettunen et al. 2019
- More information:
 - EG STAR paper 2019: A Survey on Gradient-Domain Rendering, Hua et al. 2019